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**Temporal and Spatial Variation
in Low Tide Counts of
Wildfowl and Waders on
Selected Estuaries in Great Britain**

by

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1 EXECUTIVE SUMMARY

The objectives of this report are two-fold:

- 1. To describe temporal changes in the low water distribution of waterfowl on selected estuaries in Britain based on counts taken during four months (November-February) of 1992/93.**
- 2. To define a practical count methodology which minimises bias and maximises the precision of low water count data.**

Long-term monitoring schemes are a compromise between a statistical ideal and what is practical, possible and cost-effective. To achieve the objectives outlined above, it has been necessary to evaluate all aspects of the Low Tide Count methodology. There are therefore two sections to the report. The first discusses sources of error within the counting procedure. The way in which counters are asked to define their counting units is identified as an area where increased standardisation might improve count precision and repeatability. The second section discusses the sampling procedure in relation to the number of counts required per winter. The results of the analyses carried out suggest that there is higher stochastic variability in the distribution of some species than in others. The aim of the Low Tide Count Scheme is to monitor the low tide distribution of all species of wader and wildfowl using intertidal areas of estuaries. The frequency of counts should therefore be high enough to be representative of the most variable species. The question of whether more than four counts should be carried out in a particular winter could be addressed by analysing the long-term datasets collected by the BTO as part of their studies of bird distribution in relation to tidal power.

2 GENERAL INTRODUCTION

Waterfowl which overwinter in Britain occupy different habitats both within and between estuaries. Food is the critical determinant of habitat selection within estuaries and differences in prey density account for a large proportion of the variation in bird density at low water (Goss-Custard *et al.* 1993). Numerous studies have shown that waders forage efficiently by feeding in the best places and selecting the most profitable sizes of prey (Goss-Custard *et al.* 1977). As a result, waterfowl tend to concentrate where prey density and availability are relatively high and hence energy expenditure is low. The distribution of their invertebrate prey species is in turn affected by other physical characteristics of the estuary, the most important of which are salinity and several inter-related sediment static variables such as grain size and organic content. Habitat use therefore influences how species distribute themselves both between and within estuaries with species varying in their degree of preference for specific habitat features. Species segregation is determined by differences in inter-specific morphological, dietary and behavioural differences, but can also be shaped by various random or stochastic environmental events which influence prey availability.

As overall numbers of birds on an estuary increase, densities tend to reach a maximum on the most preferred areas. Resources are always limited, however, and an upper threshold density must exist. The point at which this density is reached has been defined as the carrying capacity of an area. Sutherland and Goss-Custard's (1991) definition of the term rests on the assumption that various forms of feedback from bird density to the rate at which individuals can feed (interference, prey depletion) will cause an increasing number to fail to achieve adequate intake rates as the local bird density increases. Eventually, density will reach a level at which the addition of one further bird would result in another either starving or leaving that locality to seek a better area. However, changes in feeding conditions or, in some cases, the social system, could allow even higher bird numbers to be present in a particular locality. Distinguishing between the maximum density seen in this locality and the maximum density that is possible is a fundamentally important issue when considering the effects of changes in habitat on estuarine bird populations. Detecting levels of spatial distribution and identifying causal factors is therefore an important objective of estuarine research. In addition, baseline data on the distribution of waterfowl is of practical, conservation importance. Obtaining data on a large scale requires the development of a suitable monitoring scheme.

In designing such a scheme, a distinction must be made between census methods and sampling procedure (Blondel 1985). When considering the former, one of the most important objectives must be the precision of the counts. Lack of precision may be caused by failings in the inherent experimental design, by observer variability (since most monitoring schemes rely on large numbers of volunteers) or by differences in the way the data could be interpreted. Sampling procedure, on the other hand, is largely a question of scale. Bird populations are inherently variable. The number of plots sampled and the frequency and timing of the counts must therefore be chosen to give results that are representative of the behaviour of those populations.

Britain's estuary birds have been counted since 1969 as part of the Birds of Estuaries Enquiry (BoEE). The main function of the BoEE is to monitor the absolute numbers of waterfowl present on individual estuaries. On most estuaries this involves counting the birds in roosting flocks during the two hours either side of high water, as this is the only

time that they are concentrated onto a few, relatively accessible sites. These data enable the number of each species wintering on estuaries in Britain to be calculated and changes in population levels to be monitored.

Roosts often form at considerable distances from feeding areas (Symonds *et al.* 1984) and do not necessarily, therefore, provide information on the relative importance of different parts of individual estuaries as feeding areas for intertidal waterfowl. The increasing number of threats to estuaries has highlighted the need for information relevant to understanding their potential impact, particularly when they affect only a proportion of an estuary. Such information has, in the past, often been gathered at short notice in response to knowledge of immediate threats, with little attention paid to any standardised, comparable methods of procedure.

The aim of the National Low Tide Count Scheme, initiated in 1992, is to obtain and update regularly, information on the feeding distributions of intertidal waterfowl on all the main UK estuaries. It is therefore vital to determine the optimum procedure for achieving this aim at an early stage. This report uses the results of the first winter's low tide counts to review the current methods used and, if necessary, to recommend changes in those methods for future years' monitoring.

3 VALIDATION OF COUNTING PROCEDURE

3.1 BACKGROUND

Recording feeding distributions of intertidal estuarine birds on a nationwide basis requires a project design that is both labour-intensive and time-consuming. When organising such a project, the BTO has the advantage of a large volunteer-observer network. However, because several nationwide surveys, including the BoEE, rely on these volunteers, many of them are likely already to be involved with BoEE high tide counting. Several important constraints were considered when deciding on appropriate methods for the new National Low Tide Count Scheme. The first was the total amount of time any individual could be expected to devote to recording feeding distributions and the necessity for co-ordinating any such additional recording with their BoEE commitments. A second consideration involved the difficulties of identifying at least some species of wader over long distances and of finding birds which are foraging in creeks. Walking out over mudflats at low tide is potentially dangerous and in some circumstances impossible; recommending it to volunteers was inadvisable for safety reasons, regardless of any recording problems caused by disturbance to the birds. Observer 'quality' was therefore likely to be an even more important variable than is the case for the BoEE. Reduced visibility due to snow, mist, high winds and poor light occurs regularly in winter. An appropriate level of redundancy in the counting regime adopted was therefore required. The final consideration was the timing of the tidal cycle on particular estuaries in relation to the hours of daylight available on short winter days. In almost all circumstances, counting would be restricted to weekends, limiting the possibilities for counting at some stages of the lunar cycle. The field approach suggested for the National Low Tide Count Scheme was based on methods devised and used in earlier short-term projects carried out by the BTO and included the following six points:

- 1. To count the main UK estuaries on a five-year rotating basis, using volunteer labour as far as possible.**
- 2. To focus on the winter (November - February), with one simultaneous complete count per estuary being planned for each of the four months.**
- 3. To base data collection for each estuary on pre-established subdivisions of the intertidal area.**
- 4. To use count units averaging in the range 100-250 ha intertidal area in size.**
- 5. To conduct counts during daylight around low tide on average tides (*i.e.* neither springs nor neaps).**
- 6. To record birds separately according to whether they were feeding or roosting.**

3.2 METHODS

Thirteen estuaries were covered (Fig. 3.1). Each estuary was divided into a number of sections with one person counting each section. The boundaries of each section were chosen so that, if necessary, comparisons could be made with information collected at high tide. Initially, therefore, the BoEE boundaries for each section were used. In most cases, however, these sections were too large for only one person to count and sub-divisions were necessary within the existing BoEE sectors. Once a section had been allocated to a counter, the counter was asked to further sub-divide it into areas that were termed 'mudflats' and was provided with detailed written instructions on how to do so (Appendix A). Sections were usually split into between 1 and 10 mudflats, depending on their intertidal areas, using easily recognisable features such as the change in substrate or a permanent landmark on the mud, for example a channel marker. The need for permanence of features indicating boundaries, in order to be able to repeat counts in future years, was emphasised. There were two reasons for dividing the estuary in this manner: Firstly, the sections were easier to count if they were split into mudflats and secondly, the level of detail in the data obtained for each species was that much greater.

Counts were carried out once a month on pre-determined dates from November to February inclusive. Wherever possible, all the counts on a particular estuary took place on the same day, but on a small number of occasions this was not possible. In these situations, counters were asked to count on a date as near as possible to the one originally decided. Counts were conducted during daylight, during the period between two hours before and two hours after low tide. For each mudflat, birds were recorded separately according to whether they were feeding or roosting and an accuracy code was assigned to the count. These data, along with information on visibility, disturbance level and disturbance type were entered on specially designed forms. Detailed instructions are printed on the reverse of the form (Appendix B).

3.3 RESULTS

The results of the 1992-93 National Low Tide Counts are presented in Evans (1993) in which they are summarised on a number of distribution maps and tables. For each estuary, the following measurements were calculated and presented for each mudflat:

1. Area of mudflat (ha)
2. Peak number of birds counted (all months' data combined for each species)
3. Mean number of birds counted each month (all species)
4. Density of each species (all months' data combined).

For the purposes of this report an assessment was made of the counting procedure, both in practical and scientific terms, and is discussed below.

3.4 DISCUSSION

There are two main sources of error associated with the counting procedure: experimental design and observer variability. The latter has been the subject of considerable discussion (*e.g.* Prater 1981) and will not be discussed further here. The former involves the decision to base data collection for each estuary on pre-established sub-divisions of the intertidal area. This approach has a number of advantages. Firstly, by use of standard forms, it

facilitates rapid collection and processing of information on bird numbers and eliminates analyst variability. Secondly, where appropriate, boundaries to units can reflect changes in habitat structure. Thirdly, regardless of whether unit boundaries are based on habitat structure or lines-of-sight topographical features, etc. they can be designed to be simply and unequivocally recognisable, ensuring standardisation between counts. However, the count units defined on this basis will inevitably be variable in size. The most common alternative approach in the recording of low tide bird distributions involves use of large-scale maps to plot the positions of individual birds and flocks. These maps can then be overlain with grids of standardised size or with outlines of habitat distribution, to which birds can be assigned. This has the potential advantage of considerable flexibility in interpretation, but also has severe practical disadvantages. These include the cumbersome and time consuming nature of the operation for the counter, including post-observation map interpretation and difficulties in the field assessment of distance and direction in order to pin-point flock or bird positions on the map. For cooperative work involving counters, the simplicity of the method used in the National Low Tide Count Scheme seem the preferred approach. If mudflat size distribution is plotted for a number of estuaries (Figs. 3.2 - 3.6). Fitting a curve to the distribution makes it clear that there are two aspects to this variability - mudflats are on average smaller on smaller estuaries and there is a bias towards smaller mudflat size. Given the instructions provided, these results are not surprising, but they indicate a possibility for greater standardisation within the counting procedure. Delimitation of the optimal, average size of a recording unit necessarily involves the sort of compromise chosen for the first winter's fieldwork. Subdivision into a large number of small areas potentially gives the greatest flexibility in considering the likely impact of, for example, habitat loss. This is counterbalanced both by declining precision of usage estimates, caused by stochastic variability in bird numbers, and by potential problems of accurately delimiting numerous small count areas on intertidal flats. However, by asking counters to carry out the subdivisions of the intertidal area there is at least the possibility that they are using features that most people would find easily recognisable. Limiting mudflat size to between 100 and 250 ha would have two practical problems. Firstly, if the counter is asked to do it, there are difficulties for the counter of estimating mudflat size and deciding mudflat boundary features. Secondly, the job would be extremely time-consuming for a single person should complete standardisation be required. It may be possible, however, to give greater guidance without either of these problems occurring.

4 VALIDATION OF SAMPLING PROCEDURE

4.1 BACKGROUND

Bibby *et al.* (1992) summarised the need for distribution studies in five points:

- 1. The distribution can be related to land-use.**
- 2. Many of the conservation needs of a particular species or community can be identified by investigating habitat preferences which may manifest themselves through patterns of distribution.**
- 3. The relative value of sites of conservation importance and vulnerability can be assessed with respect to their bird fauna.**
- 4. Information valuable to environmental impact assessments is provided.**
- 5. Baseline information is generated against which future changes can be assessed.**

The sampling procedures employed in any such study must therefore be sufficiently rigorous to fulfil these aims.

Waterfowl distribution within estuaries varies within and between years and during the tidal cycle (Clark *et al.* 1993). In addition there is variability between species in the extent to which distribution and numbers are likely to change with time both within and between sites. In devising a field approach for the National Low Tide Count Scheme, a balance was struck between the scientific ideal and what it was practical to achieve using a volunteer workforce. In doing so, certain basic assumptions were made which have direct bearing on the sampling procedure. These were:

- 1. One winter's coverage in five is the minimum required for detecting changes in distribution taking place between years.**
- 2. November to February is the period of the year over which numbers of birds present tend both to be highest and to remain most consistent over an extended period.**
- 3. Three good quality counts per winter is the minimum dataset on which to assess distribution patterns. The inclusion of four counts in the programme therefore provides a level of redundancy to take into account problems with weather conditions or coverage.**

Low tide counts have been carried out successfully using a volunteer workforce at more frequent intervals than those carried out as part of the National Low Tide Count Scheme. In these cases the counts were maintained for up to five consecutive years (*e.g.* Clark *et al.* 1993). However, the time investment required in achieving and maintaining this level of coverage on the 59 main British estuaries would probably prove prohibitive. Counters found the compromise of a five year repeat time appealing and the number of counts they were expected to carry out acceptable. However, their reaction may have proved similar if, for example four monthly counts every three years had been proposed. In addition, there is no dataset containing more than five years of low tide counts with which to assess the validity of this first assumption. Leaving aside the question of the value of collecting data during passage periods, the aim of this section of the report is to assess the validity of carrying out four monthly counts on each estuary taking part in the National Low Tide Count Scheme.

4.2 METHODS

Data from five estuaries, NW Silent, Portsmouth Harbour, Swale, Forth and Wigtown Bay were used in the analysis. These were estuaries that were both well-spread geographically and on which complete coverage had been achieved.

Initially a community approach was adopted. The following comparisons were made using Spearman rank correlations:

- 1 A between-month comparison of overall bird numbers (all species combined) and overall bird density on each mudflat.
2. A comparison of early winter (November and December combined) and late winter counts (January and February combined).
3. A comparison of each count with all counts combined using both numbers and density.

It is, however, known that there is variability between species in the extent to which low tide distribution and numbers are likely to change with time within sites. A second, species-specific approach was therefore adopted which investigated this variability. A between-month comparison of the density of each species on each mudflat was made using Spearman rank correlation. In addition the percentage of significant results obtained for each species was calculated, in order to provide some measure of the differences in the predictability of different species' distribution within estuaries. The percentage of significant results obtained for each between-month comparison was calculated to give a measure of the comparability of each combination of counts.

Scatterplot matrices (SPLoMs) were plotted using the densities of four widespread species (Shelduck, Oystercatcher, Dunlin, Curlew) on one estuary, the Swale, to illustrate differences in their relative distribution. If similar densities of birds were present on each mudflat then plotting the results of one count against another should have a gradient close to 1.

4.3 RESULTS

The results of the analyses carried out for all species combined are presented in Tables 4.1 - 4.7. The relative importance of each mudflat was similar between consecutive and non-consecutive counts both when absolute numbers and when density were considered (Tables 4.1 & 4.2). Similarly, the relative importance of the mudflats did not change significantly between early and late winter counts on any of the estuaries considered (Tables 4.3 & 4.4). In addition, when the results of any single count were correlated with those of all counts combined the results were significant (Tables 4.5 & 4.6). Interestingly, the percentage of significant results obtained for each correlation between counts was similar for all combinations (Table 4.7).

The results of the analyses carried out for each species separately are presented in Tables 4.8 - 4.14. The percentage of significant results obtained varied between species (Table 4.15). The highest percentage was achieved for Grey Plover, the lowest for Curlew.

Scatterplot matrices are presented for Shelduck, Dunlin, Oystercatcher and Curlew on the Swale (Figs. 4.1 - 4.4). In all four cases, adjacent counts tended to have a gradient closer to one than counts two or three months apart. This suggests that there are changes in the numbers and distribution of these four species over the period from November to

February.

4.4 DISCUSSION

The purpose of low tide counts is to provide a standardised and reliable estimate of the relative importance of different parts of individual estuaries to waterfowl. Counting them at low tide provides a measure of the distribution of each species. Repeat counts provide a measure of the temporal variation in distribution and numbers.

A considerable number of variables can influence the numbers of intertidally feeding birds present on an estuary and the manner in which they distribute themselves. The effects of some of the more predictable environmental variables have been minimised by standardisation of the counting procedure. Two other sources of variability, however, will influence the choice of sampling procedure. Firstly, episodic events can suddenly and unpredictably affect bird feeding distribution in the dynamic estuarine environment. It has, for example, been demonstrated that large-scale changes in bird distribution occur after gales which cause soft sediment removal from exposed mudflats (Clark 1989). Secondly, even when sources of environmental variability have been controlled for, the distribution of foraging intertidal birds will vary stochastically with time. The relative importance of this source of variability in affecting estimates of feeding distribution derived from sample observations is likely to vary between species, being less important for those in which individuals tend towards being relatively solitary and sedentary, *e.g.* Redshank, than for those which tend to be more mobile and aggregated, *e.g.* Knot. These factors are evident in the results of the analysis of the first winter's National Low Tide Count data.

If there is high stochastic variability then, regardless of accuracy of counting, assessment of spatial usage patterns will be imprecise, *i.e.* have poor repeatability. This problem can be counteracted either by increasing the frequency of recording or by increasing the areas of the recording sectors into which the estuary is divided. The first section of this report considered mudflat size. It was clear that the average mudflat size chosen by counters was smaller than the recommended 100 - 250 ha. Future instructions to counters should therefore include more guidance in this respect. This second section of the report demonstrated the differing levels of variability in distribution between species. Since the aim of the Low Tide Count Scheme is to record the distribution of all species of wader and wildfowl using the intertidal area, it is logical to adopt the number of counts required to give the most accurate representation of the distribution of the largest number of species possible. In this respect, given the practical considerations discussed in this report, the number of counts adopted for the first winter's fieldwork seems likely to be the minimum required. The need for more frequent counts is a question that could be addressed by carrying out further analyses on the longer-term datasets that the BTO has collected during its studies of bird distribution in relation to tidal power.

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| SPECIES | COUNT | DEC | JAN | FEB |
|--------------------|-------|----------|----------|----------|
| INNER FORTH | NOV | 0.664*** | 0.602*** | 0.642*** |
| | DEC | - | 0.614*** | 0.664*** |
| | JAN | - | - | 0.657*** |
| OUTER FORTH SOUTH | NOV | 0.725*** | 0.651*** | 0.647*** |
| | DEC | - | 0.691*** | 0.686*** |
| | JAN | - | - | 0.696*** |
| OUTER FORTH NORTH | NOV | 0.764*** | 0.718*** | 0.677*** |
| | DEC | - | 0.699*** | 0.690*** |
| | JAN | - | - | 0.689*** |
| NW SOLENT | NOV | 0.582*** | 0.595*** | 0.573*** |
| | DEC | - | 0.514*** | 0.629*** |
| | JAN | - | - | 0.580*** |
| PORTSMOUTH HARBOUR | NOV | 0.756*** | 0.663*** | 0.693*** |
| | DEC | - | 0.774*** | 0.695*** |
| | JAN | - | - | 0.707*** |
| SWALE | NOV | 0.707*** | 0.666*** | 0.663*** |
| | DEC | - | 0.715*** | 0.691*** |
| | JAN | - | - | 0.700*** |
| WIGTOWN BAY | NOV | 0.500*** | 0.429*** | 0.489*** |
| | DEC | - | 0.550*** | 0.517*** |
| | JAN | - | - | 0.484*** |

Table 4.1 Spearman correlation coefficients (r_s) between the numbers of all species combined in four monthly counts of mudflats within each of five estuaries during the 1992/93 winter (*= $p < 0.05$, **= $p < 0.01$, ***= $p < 0.001$).

| SPECIES | COUNT | DEC | JAN | FEB |
|--------------------|-------|----------|----------|----------|
| INNER FORTH | NOV | 0.621** | 0.451* | 0.647*** |
| | DEC | - | 0.223 | 0.603** |
| | JAN | - | - | 0.581** |
| OUTER FORTH SOUTH | NOV | 0.840*** | 0.672*** | 0.690*** |
| | DEC | - | 0.676*** | 0.792*** |
| | JAN | - | - | 0.601*** |
| OUTER FORTH NORTH | NOV | 0.631*** | 0.648*** | 0.580*** |
| | DEC | - | 0.677*** | 0.662*** |
| | JAN | - | - | 0.593*** |
| NW SOLENT | NOV | 0.593* | 0.721** | 0.536* |
| | DEC | - | 0.615** | 0.760*** |
| | JAN | - | - | 0.701** |
| PORTSMOUTH HARBOUR | NOV | 0.600* | 0.424 | 0.664* |
| | DEC | - | 0.396 | 0.667* |
| | JAN | - | - | 0.464 |
| SWALE | NOV | 0.801*** | 0.414** | 0.717*** |
| | DEC | - | 0.420** | 0.686*** |
| | JAN | - | - | 0.551*** |
| WIGTOWN BAY | NOV | 0.619** | 0.486* | 0.589** |
| | DEC | - | 0.361 | 0.343 |
| | JAN | - | - | 0.439* |

Table 4.2 Spearman correlation coefficients (r_s) between the densities of all species combined in four monthly counts of mudflats within each of five estuaries during the 1992/93 winter (*= $p < 0.05$, **= $p < 0.01$, ***= $p < 0.001$).

| ESTUARY | r_s |
|---------------------------|-------------------------|
| INNER FORTH | 0.669*** |
| OUTER FORTH SOUTH | 0.736*** |
| OUTER FORTH NORTH | 0.765*** |
| NW SOLENT | 0.715*** |
| PORTSMOUTH HARBOUR | 0.755*** |
| SWALE | 0.748*** |
| WIGTOWN BAY | 0.586*** |

Table 4.3 Spearman correlation coefficients (r_s) between the numbers of all species combined in the two early winter counts (November & December) and the two late winter counts (January & February) within each of five estuaries during the 1992/93 winter (*= $p<0.05$, **= $p<0.01$, *= $p<0.001$).**

| ESTUARY | r_s |
|---------------------------|-------------------------|
| INNER FORTH | 0.583*** |
| OUTER FORTH SOUTH | 0.775*** |
| OUTER SOUTH NORTH | 0.782*** |
| NW SOLENT | 0.761*** |
| PORTSMOUTH HARBOUR | 0.567 |
| SWALE | 0.728*** |
| WIGTOWN BAY | 0.602** |

Table 4.4 Spearman correlation coefficients (r_s) between the density of all species combined in the two early winter counts (November & December) and the two late winter counts (January & February) within each of five estuaries during the 1992/93 winter (*= $p < 0.05$, **= $p < 0.01$, *= $p < 0.001$).**

| ESTUARY | COUNT | NOV + DEC + JAN + FEB |
|--------------------|-------|-----------------------|
| INNER FORTH | NOV | 0.752*** |
| | DEC | 0.773*** |
| | JAN | 0.789*** |
| | FEB | 0.753*** |
| OUTER FORTH SOUTH | NOV | 0.772*** |
| | DEC | 0.789*** |
| | JAN | 0.786*** |
| | FEB | 0.833*** |
| OUTER FORTH NORTH | NOV | 0.802*** |
| | DEC | 0.834*** |
| | JAN | 0.831*** |
| | FEB | 0.806*** |
| NW SOLENT | NOV | 0.712*** |
| | DEC | 0.686*** |
| | JAN | 0.740*** |
| | FEB | 0.715*** |
| PORTSMOUTH HARBOUR | NOV | 0.873*** |
| | DEC | 0.838*** |
| | JAN | 0.776*** |
| | FEB | 0.791*** |
| SWALE | NOV | 0.806*** |
| | DEC | 0.782*** |
| | JAN | 0.782*** |
| | FEB | 0.837*** |
| WIGTOWN BAY | NOV | 0.657*** |
| | DEC | 0.721*** |
| | JAN | 0.692*** |
| | FEB | 0.586*** |

Table 4.5 Spearman correlation coefficients (r_s) between the numbers of all species combined in each monthly count and those in all monthly counts combined for each of the five estuaries during the 1992/93 winter (*= $p < 0.05$, **= $p < 0.01$, ***= $p < 0.001$).

| ESTUARY | COUNT | NOV + DEC + JAN + FEB |
|-----------------------------|-------|-----------------------|
| INNER FORTH (n=20) | NOV | 0.848*** |
| | DEC | 0.681** |
| | JAN | 0.776*** |
| | FEB | 0.794*** |
| OUTER FORTH SOUTH (n=54) | NOV | 0.910** |
| | DEC | 0.883*** |
| | JAN | 0.855** |
| | FEB | 0.828** |
| OUTER FORTH NORTH (n=27) | NOV | 0.723*** |
| | DEC | 0.903*** |
| | JAN | 0.847*** |
| | FEB | 0.626*** |
| NW SOLENT (n=15) | NOV | 0.814*** |
| | DEC | 0.836*** |
| | JAN | 0.857*** |
| | FEB | 0.896*** |
| PORTSMOUTH HARBOUR (n=9) | NOV | 0.683* |
| | DEC | 0.917*** |
| | JAN | 0.767* |
| | FEB | 0.800** |
| SWALE (n=44) | NOV | 0.847*** |
| | DEC | 0.851*** |
| | JAN | 0.696*** |
| | FEB | 0.875*** |
| WIGTOWN BAY (n=23) | NOV | 0.878*** |
| | DEC | 0.683*** |
| | JAN | 0.668*** |
| | FEB | 0.751*** |

Table 4.6 Spearman correlation coefficients (r_s) between the densities of all species combined in each monthly count and those in all monthly counts combined for each of the five estuaries during the 1992/93 winter (*= $p < 0.05$, **= $p < 0.01$, ***= $p < 0.001$).

| COUNT | DEC | JAN | FEB |
|-------|------|------|------|
| NOV | 71.3 | 60.9 | 61.1 |
| DEC | - | 63.9 | 61.7 |
| JAN | - | - | 63.0 |

Table 4.7 The percentage of significant results obtained for the spearman rank correlations for each combination of months.

| SPECIES | COUNT | DEC | JAN | FEB |
|---------------------|-------|----------|----------|----------|
| PINK-FOOTED GOOSE | NOV | -0.050 | 0.549** | . |
| | DEC | - | -0.072 | . |
| | JAN | - | - | . |
| SHELDUCK | NOV | 0.766*** | 0.735*** | 0.587*** |
| | DEC | - | 0.572** | 0.758*** |
| | JAN | - | - | 0.492* |
| WIGEON | NOV | 0.712*** | 0.421* | 0.583** |
| | DEC | - | 0.635** | 0.800*** |
| | JAN | - | - | 0.440* |
| TEAL | NOV | 0.356 | 0.461 | 0.663*** |
| | DEC | - | 0.449* | 0.462* |
| | JAN | - | - | 0.672*** |
| MALLARD | NOV | 0.366 | 0.634** | 0.340 |
| | DEC | - | 0.260 | 0.384 |
| | JAN | - | - | 0.746*** |
| PINTAIL | NOV | 0.398 | 0.723*** | 0.724*** |
| | DEC | - | 0.606** | 0.608** |
| | JAN | - | - | 1.000*** |
| OYSTERCATCHER | NOV | 0.913*** | 0.894*** | 0.564** |
| | DEC | - | 0.863*** | 0.654** |
| | JAN | - | - | 0.650** |
| RINGED PLOVER | NOV | . | . | . |
| | DEC | - | 0.725 | 0.290 |
| | JAN | - | - | 0.486* |
| GOLDEN PLOVER | NOV | -0.196 | -0.082 | -0.086 |
| | DEC | - | 0.539 | -0.114 |
| | JAN | - | - | -0.048 |
| GREY PLOVER | NOV | . | . | . |
| | DEC | - | 0.725 | . |
| | JAN | - | - | . |
| LAPWING | NOV | 0.479* | 0.489* | 0.356 |
| | DEC | - | 0.232 | 0.786*** |
| | JAN | - | - | 0.471* |
| KNOT | NOV | 0.669*** | -0.119 | 0.291 |
| | DEC | - | 0.493* | 0.376 |
| | JAN | - | - | 0.411 |
| DUNLIN | NOV | 0.504* | 0.510* | 0.588** |
| | DEC | - | 0.511* | 0.545* |
| | JAN | - | - | 0.695*** |
| BLACK-TAILED GODWIT | NOV | 0.725*** | . | -0.100 |
| | DEC | - | . | -0.076 |
| | JAN | - | - | . |
| BAR-TAILED GODWIT | NOV | 0.523* | 0.245*** | -0.230 |
| | DEC | - | 0.377*** | 0.157 |
| | JAN | - | - | 0.424* |
| CURLEW | NOV | 0.519* | 0.656*** | -0.207 |
| | DEC | - | 0.824*** | 0.229 |
| | JAN | - | - | 0.186 |
| REDSHANK | NOV | 0.645** | 0.583** | 0.638** |
| | DEC | - | 0.718*** | 0.573** |
| | JAN | - | - | 0.839*** |
| TURNSTONE | NOV | 0.179 | 0.333 | 0.336 |
| | DEC | - | 0.487* | 0.455* |
| | JAN | - | - | 0.611** |

Table 4.8 Spearman correlation coefficients (r_s) between the density of each species in four monthly counts of mudflats within the Inner Forth during the 1992/93 winter (*= $p < 0.05$, **= $p < 0.01$, ***= $p < 0.001$).

| SPECIES | COUNT | DEC | JAN | FEB |
|-------------------|-------|----------|----------|----------|
| SHELDUCK | NOV | . | . | . |
| | DEC | - | 1.000*** | 1.000*** |
| | JAN | - | - | 0.520*** |
| WIGEON | NOV | 1.000*** | 0.718*** | 0.675*** |
| | DEC | - | 0.720*** | 1.000*** |
| | JAN | - | - | 0.472** |
| TEAL | NOV | . | 1.000*** | -0.029 |
| | DEC | - | . | . |
| | JAN | - | - | -0.038 |
| MALLARD | NOV | 0.678*** | 0.676*** | 0.628*** |
| | DEC | - | 0.556** | 0.639*** |
| | JAN | - | - | 0.719*** |
| OYSTERCATCHER | NOV | 0.743*** | 0.743*** | 0.650*** |
| | DEC | - | 0.717*** | 0.513** |
| | JAN | - | - | 0.558*** |
| RINGED PLOVER | NOV | 0.634*** | 0.629*** | 0.641*** |
| | DEC | - | 0.379* | 0.415* |
| | JAN | - | - | 0.522*** |
| GOLDEN PLOVER | NOV | 0.450* | -0.044 | -0.030 |
| | DEC | - | 0.307 | 0.232 |
| | JAN | - | - | 0.357* |
| GREY PLOVER | NOV | 0.763*** | 0.515** | 0.773*** |
| | DEC | - | 0.843*** | 0.522** |
| | JAN | - | - | 0.493** |
| LAPWING | NOV | 0.467* | 0.324 | 0.016 |
| | DEC | - | 0.498** | 0.348 |
| | JAN | - | - | 0.553*** |
| KNOT | NOV | 0.365 | 0.890*** | 0.272 |
| | DEC | - | 0.558** | 0.294 |
| | JAN | - | - | 0.644*** |
| SANDERLING | NOV | 0.599*** | 0.595*** | 0.656*** |
| | DEC | - | 1.000*** | 0.598*** |
| | JAN | - | - | 0.593*** |
| PURPLE SANDPIPER | NOV | 0.667*** | 1.000*** | 0.577*** |
| | DEC | - | 0.667*** | 0.353 |
| | JAN | - | - | 0.577*** |
| DUNLIN | NOV | 0.928*** | 0.709*** | 0.713*** |
| | DEC | - | 0.648*** | 0.480** |
| | JAN | - | - | 0.740*** |
| BAR-TAILED GODWIT | NOV | 0.761*** | 0.814*** | 0.811*** |
| | DEC | - | 0.850*** | 0.637*** |
| | JAN | - | - | 0.589*** |
| CURLEW | NOV | 0.459* | 0.348* | 0.293 |
| | DEC | - | 0.219 | 0.290 |
| | JAN | - | - | 0.442** |
| REDSHANK | NOV | 0.707* | 0.461** | 0.608*** |
| | DEC | - | 0.696*** | 0.630*** |
| | JAN | - | - | 0.537*** |
| TURNSTONE | NOV | 0.319 | 0.187 | 0.251 |
| | DEC | - | 0.562** | 0.655*** |
| | JAN | - | - | 0.510*** |

Table 4.9 Spearman correlation coefficients (r_s) between the density of each species in four monthly counts for mudflats within the Outer Forth North during the 1992/93 winter (*= $p<0.05$, **= $p<0.01$, *= $p<0.001$).**

| SPECIES | COUNT | DEC | JAN | FEB |
|-------------------|-------|----------|----------|----------|
| SHELDUCK | NOV | 0.350** | 0.617*** | 0.703*** |
| | DEC | - | 0.411** | 0.301* |
| | JAN | - | - | 0.805*** |
| WIGEON | NOV | 0.771*** | 0.496*** | 0.482*** |
| | DEC | - | 0.673*** | 0.635*** |
| | JAN | - | - | 0.396** |
| TEAL | NOV | . | . | . |
| | DEC | - | -0.019 | -0.027 |
| | JAN | - | - | 0.691*** |
| MALLARD | NOV | 0.739*** | 0.515*** | 0.337** |
| | DEC | - | 0.659*** | 0.501*** |
| | JAN | - | - | 0.485*** |
| OYSTERCATCHER | NOV | 0.803*** | 0.789*** | 0.708*** |
| | DEC | - | 0.742*** | 0.770*** |
| | JAN | - | - | 0.646*** |
| RINGED PLOVER | NOV | 0.831*** | 0.532*** | 0.555*** |
| | DEC | - | 0.581*** | 0.361** |
| | JAN | - | - | 0.582*** |
| GOLDEN PLOVER | NOV | 0.553*** | 0.310* | 0.339** |
| | DEC | - | 0.245 | 0.645*** |
| | JAN | - | - | -0.075 |
| GREY PLOVER | NOV | 0.352** | 0.694*** | 0.638*** |
| | DEC | - | 0.420** | 0.757*** |
| | JAN | - | - | 0.607*** |
| LAPWING | NOV | 0.662*** | 0.570*** | 0.452*** |
| | DEC | - | 0.816*** | 0.380** |
| | JAN | - | - | 0.504*** |
| KNOT | NOV | 0.537*** | 0.554*** | 0.487*** |
| | DEC | - | 0.370** | 0.317* |
| | JAN | - | - | 0.457*** |
| SANDERLING | NOV | -0.019 | -0.022 | 0.298* |
| | DEC | - | -0.019 | 0.456*** |
| | JAN | - | - | 0.433*** |
| PURPLE SANDPIPER | NOV | 0.123 | 0.142 | 0.142 |
| | DEC | - | 0.749*** | 0.885*** |
| | JAN | - | - | 0.718*** |
| DUNLIN | NOV | 0.767*** | 0.406*** | 0.553*** |
| | DEC | - | 0.393** | 0.463*** |
| | JAN | - | - | 0.309* |
| BAR-TAILED GODWIT | NOV | 0.789*** | 0.479*** | 0.453*** |
| | DEC | - | 0.589*** | 0.376** |
| | JAN | - | - | 0.497*** |
| CURLEW | NOV | 0.370** | 0.277* | 0.262* |
| | DEC | - | 0.309* | 0.086 |
| | JAN | - | - | 0.252* |
| REDSHANK | NOV | 0.453*** | 0.498*** | 0.387** |
| | DEC | - | 0.568*** | 0.630*** |
| | JAN | - | - | 0.433*** |
| TURNSTONE | NOV | 0.634*** | 0.606*** | 0.582*** |
| | DEC | - | 0.715*** | 0.662*** |
| | JAN | - | - | 0.699*** |

Table 4.10 Spearman correlation coefficients (r_s) between the density of each species in four monthly counts for mudflats within the Outer Forth South during the 1992/93 winter (*= $p<0.05$, **= $p<0.01$, *= $p<0.001$).**

| SPECIES | COUNT | DEC | JAN | FEB |
|---------------------|-------|----------|----------|----------|
| BRENT GOOSE | NOV | 0.283 | 0.360 | 0.361 |
| | DEC | - | 0.172 | 0.328 |
| | JAN | - | - | 0.460 |
| SHELDUCK | NOV | 0.531* | 0.391 | 0.489 |
| | DEC | - | 0.198 | 0.469 |
| | JAN | - | - | 0.626** |
| WIGEON | NOV | 0.619* | 0.781*** | 0.003 |
| | DEC | - | 0.655** | 0.182 |
| | JAN | - | - | 0.434 |
| TEAL | NOV | 0.488 | 0.255 | 0.382 |
| | DEC | - | -0.200 | 0.460 |
| | JAN | - | - | 0.270 |
| MALLARD | NOV | 0.463 | 0.238 | 0.879*** |
| | DEC | - | -0.133 | 0.467 |
| | JAN | - | - | 0.249 |
| OYSTERCATCHER | NOV | 0.172 | 0.090 | 0.500 |
| | DEC | - | 0.010 | 0.000 |
| | JAN | - | - | 0.593** |
| RINGED PLOVER | NOV | 0.627* | -0.133 | 0.556* |
| | DEC | - | -0.200 | 0.270 |
| | JAN | - | - | 0.604* |
| GOLDEN PLOVER | NOV | 1.000*** | 1.000*** | 0.516* |
| | DEC | - | 1.000*** | 0.507* |
| | JAN | - | - | 0.503* |
| GREY PLOVER | NOV | 0.561* | 0.523* | 0.588* |
| | DEC | - | 0.537* | 0.917*** |
| | JAN | - | - | 0.613** |
| LAPWING | NOV | 0.704* | 0.799*** | 0.663** |
| | DEC | - | 0.603* | 0.526* |
| | JAN | - | - | 0.697** |
| KNOT | NOV | . | 0.627* | 0.619* |
| | DEC | - | . | . |
| | JAN | - | - | 0.780* |
| DUNLIN | NOV | 0.613* | 0.219 | 0.502 |
| | DEC | - | 0.630** | 0.661** |
| | JAN | - | - | 0.456 |
| BLACK-TAILED GODWIT | NOV | 1.000*** | 1.000*** | 0.759*** |
| | DEC | - | 1.000*** | 0.753*** |
| | JAN | - | - | 0.750*** |
| BAR-TAILED GODWIT | NOV | . | . | . |
| | DEC | - | 0.480 | 0.380 |
| | JAN | - | - | 0.398 |
| CURLEW | NOV | 0.525* | 0.195 | 0.604* |
| | DEC | - | -0.001 | 0.336 |
| | JAN | - | - | -0.021 |
| REDSHANK | NOV | 0.260 | 0.579* | 0.508 |
| | DEC | - | 0.296 | 0.539* |
| | JAN | - | - | 0.429 |
| TURNSTONE | NOV | 0.589* | 0.752** | 0.553* |
| | DEC | - | 0.447 | 0.648** |
| | JAN | - | - | 0.156 |

Table 4.11 Spearman correlation coefficients (r_s) between the density of each species in four monthly counts for mudflats within the NW Solent during the 1992/93 winter (*= $p<0.05$, **= $p<0.01$, *= $p<0.001$).**

| SPECIES | COUNT | DEC | JAN | FEB |
|---------------------|-------|----------|----------|---------|
| BRENT GOOSE | NOV | 0.528 | 0.121 | -0.136 |
| | DEC | - | 0.593** | 0.167 |
| | JAN | - | - | 0.497 |
| SHELDUCK | NOV | 0.734** | 0.633** | 0.579 |
| | DEC | - | 0.658** | 0.700* |
| | JAN | - | - | 0.672* |
| WIGEON | NOV | 0.323 | . | . |
| | DEC | - | . | . |
| | JAN | - | - | . |
| TEAL | NOV | -0.113 | . | . |
| | DEC | - | . | . |
| | JAN | - | - | . |
| MALLARD | NOV | 0.754** | 0.486 | -0.91 |
| | DEC | - | 0.170 | -0.300 |
| | JAN | - | - | -0.100 |
| OYSTERCATCHER | NOV | 0.876*** | 0.253 | 0.510 |
| | DEC | - | 0.539* | 0.729* |
| | JAN | - | - | 0.566 |
| RINGED PLOVER | NOV | 0.774** | 0.403 | . |
| | DEC | - | 0.329 | . |
| | JAN | - | - | . |
| GREY PLOVER | NOV | 0.295 | 0.492 | 0.784** |
| | DEC | - | 0.230 | 0.443 |
| | JAN | - | - | 0.371 |
| LAPWING | NOV | 0.471 | 0.592** | 0.500 |
| | DEC | - | 0.812*** | 0.750* |
| | JAN | - | - | 0.500 |
| DUNLIN | NOV | 0.820*** | 0.393 | 0.702* |
| | DEC | - | 0.404 | 0.409 |
| | JAN | - | - | 0.641* |
| BLACK-TAILED GODWIT | NOV | 0.444 | 0.523* | . |
| | DEC | - | 0.583** | . |
| | JAN | - | - | . |
| BAR-TAILED GODWIT | NOV | -0.077 | . | -0.100 |
| | DEC | - | -0.050 | -0.125 |
| | JAN | - | - | . |
| CURLEW | NOV | 0.593* | 0.264 | 0.564 |
| | DEC | - | 0.295 | 0.745* |
| | JAN | - | - | 0.616* |
| REDSHANK | NOV | 0.815*** | 0.387 | 0.545 |
| | DEC | - | 0.302 | 0.783* |
| | JAN | - | - | 0.264 |
| TURNSTONE | NOV | -0.067 | 0.482 | 0.259 |
| | DEC | - | 0.427 | 0.380 |
| | JAN | - | - | 0.500 |

Table 4.12 Spearman correlation coefficients (r_s) between the density of each species in four monthly counts for mudflats within the Portsmouth Harbour during the 1992/93 winter (*= $p<0.05$, **= $p<0.01$, *= $p<0.001$).**

| SPECIES | COUNT | DEC | JAN | FEB |
|---------------------|-------|----------|----------|----------|
| BRENT GOOSE | NOV | -0.049 | -0.093 | -0.100 |
| | DEC | - | 0.100 | 0.250 |
| | JAN | - | - | 0.491*** |
| SHELDUCK | NOV | 0.660*** | 0.715*** | 0.752*** |
| | DEC | - | 0.898*** | 0.834*** |
| | JAN | - | - | 0.813*** |
| WIGEON | NOV | -0.028 | -0.030 | -0.047 |
| | DEC | - | 0.807*** | 0.544*** |
| | JAN | - | - | 0.672*** |
| TEAL | NOV | -0.028 | -0.043 | 0.433** |
| | DEC | - | -0.029 | 0.361* |
| | JAN | - | - | 0.270 |
| MALLARD | NOV | 0.527*** | 0.704*** | 0.422** |
| | DEC | - | 0.810*** | 0.604*** |
| | JAN | - | - | 0.621*** |
| PINTAIL | NOV | 0.480*** | -0.030 | 0.384*** |
| | DEC | - | 0.714*** | 0.825*** |
| | JAN | - | - | 0.591*** |
| OYSTERCATCHER | NOV | 0.761*** | 0.656*** | 0.573*** |
| | DEC | - | 0.786*** | 0.717*** |
| | JAN | - | - | 0.649*** |
| AVOCET | NOV | . | -0.021 | -0.032 |
| | DEC | - | . | . |
| | JAN | - | - | 0.682*** |
| RINGED PLOVER | NOV | 0.376** | 0.193 | 0.106 |
| | DEC | - | 0.405** | 0.496*** |
| | JAN | - | - | 0.114 |
| GOLDEN PLOVER | NOV | -0.081 | 0.330* | -0.046 |
| | DEC | - | -0.098 | 0.306* |
| | JAN | - | - | -0.069 |
| GREY PLOVER | NOV | 0.561*** | 0.770*** | 0.694*** |
| | DEC | - | 0.418** | 0.451** |
| | JAN | - | - | 0.744*** |
| LAPWING | NOV | 0.510*** | 0.071 | 0.291 |
| | DEC | - | 0.487*** | 0.589*** |
| | JAN | - | - | 0.428** |
| KNOT | NOV | 0.659*** | 0.531*** | 0.405** |
| | DEC | - | 0.594*** | 0.560*** |
| | JAN | - | - | 0.418** |
| DUNLIN | NOV | 0.577*** | 0.419** | 0.598*** |
| | DEC | - | 0.619*** | 0.547*** |
| | JAN | - | - | 0.489*** |
| BLACK-TAILED GODWIT | NOV | 0.412*** | 0.474*** | 0.587*** |
| | DEC | - | 0.254 | 0.353* |
| | JAN | - | - | 0.413* |
| BAR-TAILED GODWIT | NOV | 0.629*** | 0.501*** | 0.468** |
| | DEC | - | 0.379** | 0.600*** |
| | JAN | - | - | 0.321* |
| CURLEW | NOV | 0.440*** | 0.497*** | 0.480*** |
| | DEC | - | 0.268 | 0.467*** |
| | JAN | - | - | 0.371 |
| REDSHANK | NOV | 0.502*** | 0.570*** | 0.545*** |
| | DEC | - | 0.631*** | 0.665*** |
| | JAN | - | - | 0.656*** |
| TURNSTONE | NOV | 0.491*** | 0.503*** | 0.420** |
| | DEC | - | 0.624*** | 0.528*** |
| | JAN | - | - | 0.510*** |

Table 4.13 Spearman correlation coefficients (r_s) between the density of each species in four monthly counts for mudflats within the Swale during the 1992/93 winter (*= $p<0.05$, **= $p<0.01$, *= $p<0.001$).**

| SPECIES | COUNT | DEC | JAN | FEB |
|-------------------|-------|----------|----------|----------|
| PINK-FOOTED GOOSE | NOV | . | . | . |
| | DEC | - | -0.119 | -0.200 |
| | JAN | - | - | 0.071 |
| SHELDUCK | NOV | 0.207 | 0.181 | 0.496* |
| | DEC | - | 0.180 | 0.471* |
| | JAN | - | - | 0.202 |
| WIGEON | NOV | 0.451* | 0.650*** | 0.262 |
| | DEC | - | 0.207 | 0.317 |
| | JAN | - | - | 0.336 |
| TEAL | NOV | 1.000*** | 0.723*** | 0.722*** |
| | DEC | - | 0.723*** | 0.723*** |
| | JAN | - | - | 0.500* |
| MALLARD | NOV | 0.300 | 0.292 | 1.000*** |
| | DEC | - | 0.631** | 0.300 |
| | JAN | - | - | 0.292 |
| PINTAIL | NOV | 0.657*** | 0.535** | 1.000*** |
| | DEC | - | 0.700*** | 0.657*** |
| | JAN | - | - | 0.535** |
| OYSTERCATCHER | NOV | 0.832*** | 0.807*** | 0.662*** |
| | DEC | - | 0.890*** | 0.606** |
| | JAN | - | - | 0.612** |
| RINGED PLOVER | NOV | . | -0.066 | . |
| | DEC | - | . | . |
| | JAN | - | - | . |
| GOLDEN PLOVER | NOV | 0.450* | -0.066 | 0.358 |
| | DEC | - | -0.066 | 0.399 |
| | JAN | - | - | -0.082 |
| LAPWING | NOV | 0.357 | -0.176 | 0.079 |
| | DEC | - | 0.263 | 0.202 |
| | JAN | - | - | -0.295 |
| KNOT | NOV | . | . | . |
| | DEC | - | . | -0.066 |
| | JAN | - | - | . |
| DUNLIN | NOV | 0.791*** | 0.125 | -0.035 |
| | DEC | - | 0.301 | 0.117 |
| | JAN | - | - | 0.500* |
| BAR-TAILED GODWIT | NOV | -0.082 | -0.082 | 0.464* |
| | DEC | - | 1.000*** | -0.045 |
| | JAN | - | - | -0.045 |
| CURLEW | NOV | 0.232 | 0.399 | 0.355 |
| | DEC | - | 0.010 | 0.196 |
| | JAN | - | - | 0.275 |
| REDSHANK | NOV | -0.033 | 0.184 | 0.487* |
| | DEC | - | 0.507* | 0.305 |
| | JAN | - | - | 0.307 |
| TURNSTONE | NOV | 0.604** | 0.604** | . |
| | DEC | - | 1.000*** | . |
| | JAN | - | - | . |

Table 4.14 Spearman correlation coefficients (r_s) between the density of each species in four monthly counts for mudflats within the Wigtown Bay during the 1992/93 winter (*= $p<0.05$, **= $p<0.01$, *= $p<0.001$).**

| SPECIES | % |
|----------------------------|-------------|
| Grey Plover | 83.9 |
| Dunlin | 83.3 |
| Black-tailed Godwit | 77.8 |
| Shelduck | 76.9 |
| Oystercatcher | 76.2 |
| Redshank | 71.4 |
| Bar-tailed Godwit | 71.0 |
| Wigeon | 70.0 |
| Turnstone | 64.1 |
| Lapwing | 59.5 |
| Mallard | 59.5 |
| Ringed Plover | 56.0 |
| Golden Plover | 50.0 |
| Teal | 47.1 |
| Curlew | 45.2 |

Table 4.15 The percentage of significant results obtained for the spearman rank correlations for each species.

